

# Math 5810, Review for Exam 2, Solutions

1. a)  $\left(\frac{9}{10}\right)^5$    b)  $\frac{1}{1/10} = 10$    c)  $20 \cdot \frac{1}{10} = 2$    d)  $\binom{19}{2} \left(\frac{1}{10}\right)^3 \left(\frac{9}{10}\right)^{17}$

2. a)  $P(X \leq \frac{1}{2}) = \int_0^{\frac{1}{2}} 12x^2(1-x)dx = \int_0^{\frac{1}{2}} (12x^2 - 12x^3)dx = [4x^3 - 3x^4]_0^{\frac{1}{2}}$   
 $= 4 \cdot \left(\frac{1}{2}\right)^3 - 3 \left(\frac{1}{2}\right)^4 = \frac{1}{2} - \frac{3}{16} = \frac{5}{16}.$

b)  $P(X = \frac{1}{3}) = 0$

c)  $E(X) = \int_0^1 12x^3(1-x)dx = \int_0^1 (12x^3 - 12x^4)dx = [3x^4 - \frac{12}{5}x^5]_0^1$   
 $= 3 - \frac{12}{5} = \frac{3}{5}$

d)  $E(X^2) = \int_0^1 12x^4(1-x)dx = \int_0^1 (12x^4 - 12x^5)dx = [\frac{12}{5}x^5 - 2x^6]_0^1$   
 $= \frac{12}{5} - 2 = \frac{2}{5} \Rightarrow \text{Var}(X) = \frac{2}{5} - \left(\frac{3}{5}\right)^2 = \frac{1}{25}.$

3.  $P(X=k) = P(k \leq T < k+1) = P(T \geq k) - P(T \geq k+1) = e^{-\lambda k} - e^{-\lambda(k+1)}$   
 $= e^{-\lambda k} (1 - e^{-\lambda}) = (e^{-\lambda})^k (1 - e^{-\lambda}),$  so  $X$  has a (shifted) geometric distribution on  $\{0, 1, 2, \dots\}$  with parameter  $p = 1 - e^{-\lambda}.$

4. First, since  $\text{Range}(T) = (0, \infty), \text{Range}(Y) = (0, 1).$  For  $0 < y < 1,$   
 $F_Y(y) = P(Y \leq y) = P(e^{-T} \leq y) = P(-T \leq \log y) = P(T \geq -\log y)$   
 $= e^{-\lambda(-\log y)} = e^{\lambda \log y} = (e^{\log y})^\lambda = y^\lambda, \text{ and } f_Y(y) = F_Y'(y) = \lambda y^{\lambda-1}.$

5. Assume the weights of the hotel guests are independent of each other and all have the same distribution. Let  $X_i =$  weight of the  $i^{\text{th}}$  guest in the elevator,  $i = 1, \dots, 30,$  and let  $S_{30} := X_1 + \dots + X_{30}$  be the total weight. Then  $E(S_{30}) = (30)(150) = 4500$  and  $SD(S_{30}) = \sqrt{30(55)} = 301.25.$  By the Central Limit Theorem,

$$P(S_{30} > 5000) \approx 1 - \Phi\left(\frac{5000 - 4500}{301.25}\right) = 1 - \Phi(1.66) = 1 - .9515 = \boxed{.0485}$$

Note: Continuity correction makes little sense here (why?), but also does little harm.

6. Let  $T$  be the number of tosses until the first head shows. Then  $T \sim \text{geometric}(p).$  Let  $A =$  "A tosses the first head",  $B =$  "B tosses the first head".

Then  $P(A) = P(T \text{ is odd}) = \sum_{n=0}^{\infty} P(T = 2n+1) = \sum_{n=0}^{\infty} q^{2n} p = \frac{p}{1-q^2} = \frac{1}{1+q}$

and  $P(B) = 1 - P(A) = \frac{q}{1+q} \leq P(A),$  with equality if and only if  $q = 1$  (so  $p = 0$ ).

7. Let  $A_i =$  "the  $i^{\text{th}}$  letter is placed in the correct envelope",  $i = 1, \dots, n.$  Let  $N$  be the number of letters placed in the correct envelopes. Then  $P(A_i) = \frac{1}{n}$  for each  $i,$  and  $N = I_{A_1} + \dots + I_{A_n}.$  Thus,

$$E(N) = P(A_1) + \dots + P(A_n) = n \cdot \frac{1}{n} = 1.$$

8. Let  $N$  be the total number of misprints, and  $X$  the number of misprints found. Then  $N \sim \text{Poisson}(\mu)$ , and  $(X|N=n) \sim \text{Binomial}(n, p)$ .  
So for  $k=0, 1, 2, \dots$ ,

$$\begin{aligned}
 P(X=k) &= \sum_{n=k}^{\infty} P(X=k, N=n) = \sum_{n=k}^{\infty} P(X=k|N=n) P(N=n) \\
 &= \sum_{n=k}^{\infty} \binom{n}{k} p^k (1-p)^{n-k} e^{-\mu} \frac{\mu^n}{n!} \\
 &= e^{-\mu} \cdot p^k \cdot \frac{1}{k!} \sum_{n=k}^{\infty} \frac{(1-p)^{n-k}}{(n-k)!} \cdot \mu^n \\
 &= e^{-\mu} \frac{(\mu p)^k}{k!} \sum_{n=k}^{\infty} \frac{[\mu(1-p)]^{n-k}}{(n-k)!} \\
 &= e^{-\mu} \frac{(\mu p)^k}{k!} \cdot \sum_{j=0}^{\infty} \frac{[\mu(1-p)]^j}{j!} \\
 &= e^{-\mu} \frac{(\mu p)^k}{k!} \cdot e^{\mu(1-p)} \\
 &= e^{-\mu p} \frac{(\mu p)^k}{k!}
 \end{aligned}$$

$\therefore X \sim \text{Poisson}(\mu p)$ .