

# Math 4610, Review for Exam 3, Solutions

1. a)  $\left(\frac{9}{10}\right)^5$  b)  $\frac{1}{10} = 10$  c)  $20 \cdot \frac{1}{10} = 2$  d)  $\binom{19}{2} \left(\frac{1}{10}\right)^3 \left(\frac{9}{10}\right)^{17}$

2. Let  $T$  be the atom's lifetime and assume  $T \sim \text{Exp}(\lambda)$ . Then  $\lambda = \frac{\ln 2}{1000}$  and so  
 $P(T > t) = e^{-\lambda t} = (e^{\ln 2})^{-t/1000} = 2^{-t/1000}$ . Thus, using the memoryless property,

$$P(T \leq 1500 | T > 1000) = 1 - P(T > 1500 | T > 1000) = 1 - P(T > 500) = 1 - 2^{-500/1000} = 1 - 2^{-1/2} = \boxed{.2929}$$

3.  $E(XY) = E(X)E(Y) = \mu^2$ , and  $E[(XY)^2] = E[X^2Y^2] = E(X^2)E(Y^2)$   
 $= \{ \text{Var}(X) + (E(X))^2 \} \cdot \{ \text{Var}(Y) + (E(Y))^2 \} = (\sigma^2 + \mu^2)^2$ , so  $\text{Var}(XY) = E[(XY)^2] - [E(XY)]^2$   
 $= (\sigma^2 + \mu^2)^2 - (\mu^2)^2 = (\sigma^4 + 2\sigma^2\mu^2 + \mu^4) - \mu^4 = \sigma^4 + 2\sigma^2\mu^2 = \sigma^2(\sigma^2 + 2\mu^2)$ .

4. a)  $\int_0^1 x(1-x)^2 dx = \int_0^1 (x - 2x^2 + x^3) dx = \frac{1}{2} - \frac{2}{3} + \frac{1}{4} = \frac{1}{12} \Rightarrow \boxed{c=12}$

b)  $P(X \leq \frac{1}{2}) = \int_0^{1/2} 12x(1-x)^2 dx = [6x^2 - 8x^3 + 3x^4]_0^{1/2} = \frac{3}{2} - 1 + \frac{3}{16} = \boxed{\frac{11}{16}}$

c)  $P(X = \frac{1}{3}) = 0$ .

d)  $E(X) = \int_0^1 12x^2(1-x)^2 dx = [4x^3 - 6x^4 + \frac{12}{5}x^5]_0^1 = 4 - 6 + \frac{12}{5} = \boxed{\frac{2}{5}}$ .

e)  $E(X^2) = \int_0^1 12x^3(1-x)^2 dx = \int_0^1 (12x^3 - 24x^4 + 12x^5) dx = \frac{12}{4} - \frac{24}{5} + \frac{12}{6} = \frac{1}{5}$

$\therefore \text{Var}(X) = E(X^2) - [E(X)]^2 = \frac{1}{5} - \frac{4}{25} = \boxed{\frac{1}{25}}$ .

5. Assume the weights of the hotel guests are independent of each other and all have the same distribution. Let  $X_i =$  weight of the  $i^{\text{th}}$  guest in the elevator,  $i = 1, \dots, 30$ , and let  $S_{30} := X_1 + \dots + X_{30}$  be the total weight. Then  $E(S_{30}) = (30)(150) = 4500$  and  $SD(S_{30}) = \sqrt{30}(55) = 301.25$ . By the Central Limit Theorem,

$$P(S_{30} > 5000) \approx 1 - \Phi\left(\frac{5000 - 4500}{301.25}\right) = 1 - \Phi(1.66) = 1 - .9515 = \boxed{.0485}$$

Note: Continuity correction makes little sense here (why?), but also does little harm.

6. a) Let  $X$  be the number of brown chocolate chips in this cookie (not "per cookie", "in a cookie")

It's reasonable to assume  $X \sim \text{Ps}(4)$ , so  $P(X \leq 2) = e^{-4}(1 + 4 + \frac{4^2}{2}) = 13e^{-4} \approx \boxed{.2381}$

b) Let  $Y$  be the number of white chocolate chips in the cookie; assume  $Y \sim \text{Ps}(2)$  and  $Y$  is independent of  $X$ . Let  $Z := X + Y$  be the total number of chocolate chips in the cookie.

Then  $Z \sim \text{Ps}(6)$ , so  $P(Z \leq 5) = e^{-6}(1 + 6 + \frac{6^2}{2!} + \frac{6^3}{3!} + \frac{6^4}{4!} + \frac{6^5}{5!}) \approx \boxed{.4457}$

c) Under the same assumptions as before,  $P(X=3, Y=2 | Z=5) = \frac{P(X=3, Y=2)}{P(Z=5)}$

$$= \frac{e^{-2} \cdot \frac{2^2}{2!} \cdot e^{-4} \cdot \frac{4^2}{2!}}{e^{-6} \cdot \frac{6^5}{5!}} = \frac{5!}{2!3!} \cdot \frac{2^2 \cdot 4^2}{6^5} = \binom{5}{2} \left(\frac{1}{3}\right)^2 \left(\frac{2}{3}\right)^3 \approx \boxed{0.3292}$$

7. a)  $N = I_A + I_B$ ; b)  $E(N) = E(I_A) + E(I_B) = P(A) + P(B) = \frac{1}{3} + \frac{1}{2} = \frac{5}{6}$ .

c)  $\text{Var}(N) = \text{Var}(I_A) + \text{Var}(I_B)$  [because  $I_A, I_B$  independent!]  $= \frac{1}{3}(1-\frac{1}{3}) + \frac{1}{2}(1-\frac{1}{2}) = \frac{17}{36}$

d) Since  $A \subset B$ , if  $A$  occurs, then  $B$  occurs also. Thus,  $P(N=0) = P(B^c) = \frac{1}{2}$ ,

$P(N=1) = P(B \text{ but not } A) = P(B) - P(A) = \frac{1}{2} - \frac{1}{3} = \frac{1}{6}$ ,  $P(N=2) = P(A \text{ and } B) = P(A) = \frac{1}{3}$ .

So  $E(N^2) = 1^2 \cdot \frac{1}{6} + 2^2 \cdot \frac{1}{3} = \frac{1}{6} + \frac{4}{3} = \frac{3}{2}$ , and  $\text{Var}(N) = \frac{3}{2} - (\frac{5}{6})^2 = \frac{54}{36} - \frac{25}{36} = \frac{29}{36}$ .

8. a)  $P(A_n) = P(\text{first } n-1 \text{ draws, then } A \text{ wins}) = p_D^{n-1} p_A$ .

b)  $P(A \text{ wins overall}) = \sum_{n=1}^{\infty} P(A \text{ wins overall in } n \text{ games}) = \sum_{n=1}^{\infty} P(A_n) = \sum_{n=1}^{\infty} p_D^{n-1} p_A$   
 $= p_A \cdot \frac{1}{1-p_D} = \frac{p_A}{p_A + p_B}$ .

9. Let  $T$  be your total waiting time; then  $T \sim \text{exp}(\lambda)$  with  $\lambda = \frac{1}{10}$ .

The required probability is

$P(T \leq 12 | T > 10) = 1 - P(T > 12 | T > 10) = 1 - P(T > 2)$  (memoryless property!)  
 $= 1 - e^{-\lambda \cdot 2} = 1 - e^{-1/5} \approx .1813$ .

10. Let  $N = \text{total \# of misprints}$ ,  $X = \text{\# of misprints found}$ . Then  $P(X=k) = \sum_{n=k}^{\infty} P(X=k | N=n) P(N=n)$   
 $= \sum_{n=k}^{\infty} \binom{n}{k} p^k (1-p)^{n-k} e^{-\mu} \frac{\mu^n}{n!} = e^{-\mu} \sum_{n=k}^{\infty} \frac{p^k (1-p)^{n-k} \mu^n}{k! (n-k)!} = e^{-\mu} \frac{p^k \mu^k}{k!} \sum_{n=k}^{\infty} \frac{(1-p)^{n-k} \mu^{n-k}}{(n-k)!}$   
 $= e^{-\mu} \frac{(\mu p)^k}{k!} \sum_{l=0}^{\infty} \frac{((1-p)\mu)^l}{l!} = e^{-\mu} \frac{(\mu p)^k}{k!} e^{(1-p)\mu} = e^{-\mu p} \frac{(\mu p)^k}{k!} \therefore \boxed{X \sim P_s(\mu p)}$